

Subnational Persistence of State History: Evidence from Geolocalized Civilizations.*

C. Justin Cook
Tulane University
Murphy Institute and Dept. of Economics

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Abstract

This paper uses a novel data set that geolocalizes civilizations across the globe and human history. The data document roughly 1,450 civilizations over 500 periods between 3200 BCE and 2006 CE. Geolocalized data are then spatially aggregated to measure state history for pre-defined grids. Grid-cells with a prolonged history of containing a civilization have larger contemporary populations and increased nighttime lights. The state history data also allow for heterogeneity tests that have not been possible previously. My results show the sub-national relationship between state history and contemporary development is unaltered by country-level characteristics, suggesting an alternative mechanism is at play.

JEL-Classification: O11, O43, N10, R12

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Email: Justin Cook ccook7@tulane.edu

1 Introduction

People and economic output are not randomly distributed across the world; instead, they are tied to historic settlements, or spatially persistent. Indeed, cross-country correlation coefficients show a strong persistence in population densities across time, and this positive relationship between historic settlement patterns and contemporary populations and economic activity has been documented in a growing literature that explores the proposed stickiness of populations (and resulting output) both across and within countries.¹ This paper seeks to contribute to this literature by introducing a novel data set that synthesizes the cross- and within-country approaches. The data also allow for new tests of mechanisms that examine whether country-level factors like institutional quality are driving the estimated relationship between historic settlements and contemporary output.

In order to measure historical settlements, I use a new data set that geolocalizes the position of roughly 1,450 civilizations from 3200 BCE to the present day (c.2006). These data dynamically plot civilization borders, documenting expansions and contractions over time. The data are geolocated and not bound by contemporary country borders, which are often used as units of analysis for periods preceding the formation of these boundaries. The time and location of all civilizations are then collapsed to a spatially distributed measure that catalogs the accumulated historical presence of a civilization for a sub-national cell, creating a measure of state history similar in concept and construction to those of Bockstette et al. (2002) and Borcan et al. (2019).

I follow the empirical framework of (Henderson et al., 2018, hereafter HSSW), who create a spatial quarter degree squared grid of inhabitable land and show that a parsimonious set of environmental controls tied to agriculture and trade explain roughly half of the variation in nighttime lights. HSSW also argue that agricultural variables serve as a substitute for historical societal organization, but they are unable to directly show this because “[they] can’t observe the detailed locations of historical agglomerations[.]” This paper constructs a proxy for these historical agglomerations and tests for an association with contemporary nighttime lights and population, conditional on HSSW’s geo-climatic covariates. In other words, I take the $\approx 240\text{K}$ grid cells of HSSW, count the years a cell has been part of a civilization (i.e., state history), then test whether cells that have a prolonged exposure to civilization (or more state history) also have more contemporary economic activity—proxied by nighttime lights—and larger contemporary populations.²

¹The correlation coefficient is 0.88 between 1 CE and 1500 CE population densities and 0.62 between 1500 CE and 2000 CE densities.

²My granular spatial state history data are also not beholden to modern day country borders, which were endogenously formed by natural or colonized means that are possibly tied to current economic conditions (Borcan, Olsson, and Putterman, 2018). Civilization borders were endogenously formed too, but these historic borders do not perfectly align with current borders and can be seen as conditionally randomized

The spatial data allow for across and *within* country estimations, where within country variation allows for the explicit control of federal institutions and other unobserved country-level differences. Across all specifications I do indeed find a longer state history is associated with richer and more populated cells today, a finding inline with prior studies (discussed in detail in Section 1.1). But I also find that controls for broad country-level differences in institutional quality make little difference in the estimated relationship, suggesting the estimated effect of state history on output/population is independent of accumulated differences in institutional quality. This finding runs counter to prior studies that argue prolonged exposure to higher-order state formation leads to learning-by-doing in state capacity that manifests in country-level differences in federal institutions and ultimately, a range of socioeconomic measures today (Borcan, Olsson, and Putterman, 2018; Lagerlöf, 2016).

The effect of state history being independent of country-level factors suggests a mechanism other than contemporary institutional quality is at play. One possibility is that the persistence of populations and economic activity are due to agglomeration effects, or hard-to-measure externalities, that contribute to the persistence of cities and states, such that populations build upon historic settlements rather than move to geoclimatically identical, but new locations. Agglomeration effects are more likely in urban locations (Glaeser, 2010). Examples of this are numerous, from Mexico City in the New World to Rome, Athens, Damascus, and Luoyang in the Old World. Historic markets, or centers of trade, serve as focal points that persist over time, creating a strong link between the historical presence of the city and contemporary population/output. With this in mind, I test whether the effect of state history is stronger in urban environments, decaying as cells mover farther away from the city-center.

Baseline estimates show areas with a more prolonged state history are richer, proxied by being more well lit, and more populated today. A standard deviation increase in state history leads to a 100% increase in HSSW’s measure of nighttime lights (i.e., modified DMSP), a 110% increase in VIIRS nighttime lights, and a 140% increase in population.³ These estimated relationships are unaffected when including or excluding country FE or direct measures for institutional quality, suggesting a mechanism that is independent of

within a country. That is, when controlling for agricultural conditions and country fixed effects, the historic placement of a civilization within a country is more randomly determined than modern country borders. Indeed, the approach I take in the current paper is explicitly laid out in BOP (p. 3): “A potential alternative to using country borders could have been to divide the world into equal-sized grid cells (or “virtual countries”) and then study the history of each such cell.” Depetris-Chauvin (2015) and Maloney and Valencia Caicedo (2016) also perform similar analyses; however, each is constrained to particular time periods and locations—respectively, post-1000 CE sub-Saharan Africa and post-1492 America. My approach instead considers the global distribution of all civilizations from 3200 BCE to the present.

³These effects are large but similar to the effects of control variables.

institutional quality. To this point, the effect of state history is influenced by the proximity of a city, becoming insignificantly different than zero roughly 600 kilometers outside of an urban location. Taken together, these heterogeneity tests suggests some role for agglomeration that is independent of institutional quality.

The rest of the paper is organized as follows: Section 1.1 provides a background of prior studies linking state history to contemporary outcomes; Section 2.1 gives a descriptions of the primary state history variable and all other variables used in the paper; Section 2.2 compares the constructed state history measure to other similar measures; Section 2.3 gives the baseline empirical specification, the results of which are given in Sec. 3.1; heterogeneity tests of mechanisms are discussed in Sec. 3.2; and robustness exercises for the relationship between state history and contemporary outcomes are discussed in Sec. 3.3.

1.1 Related Literature

My measure of state history is based on a similar measure from (Bockstette, Chanda, and Putterman, 2002, hereafter BCP) and (Borcan, Olsson, and Putterman, 2018, hereafter BOP). BCP provides evidence that country-level growth in GDP per capita following decolonization in the 1960s is positively associated with state capacity, which is determined by state history. The mechanism being that longer lived states are more stable and have higher institutional quality and that this stability contributed to faster growth rates.

BOP extends this measure of state history and argues for non-linear associations between state history and state capacity. The hypothesis being that state history increases state capacity up to a point with prolonged exposure leading to over-centralization, creating a hump-shaped (inverted-U) relationship between state history and current development (Hariri, 2012; Lagerlöf, 2016). The non-linear mechanism of BOP are again primarily tied to institutional quality from longer lived states.

My approach builds off related sub-national research within Africa (Depetris-Chauvin, 2015; Michalopoulos and Papaioannou, 2013). Michalopoulos and Papaioannou use ancestral ethnic boundaries and a measure of state formation for each ethnicity to estimate the effect of higher order governance on contemporary economic output, proxied by night-time lights; indeed, they find a strong, positive, and robust relationship between ancestral state formation and contemporary development. My approach is most similar to work by Depetris-Chauvin (2015), who documents the boundaries of civilizations in sub-Saharan and connects these to modern conflict. Figure A.2 directly compares Depetris-Chauvin’s measure of state history to my own for Africa, showing considerable overlaps between the independently constructed measures. My approach extends the core idea of these papers by plotting borders for civilizations across the globe, not just Africa, and extends the time frame to cover the historic universe of human settlements.

Outside of Africa, work by (Maloney and Valencia Caicedo, 2016, hereafter MVC) uses historic censuses to estimate the persistence of populations. Within the Americas, states/provinces (from contemporary borders) that contained more people per square kilometer before the arrival of Columbus have a strong positive association with contemporary population densities and incomes. While MVC’s analysis is limited to North and South America, my approach broadens the proposed relationship, both globally and deeper into the past; however, the degree of intensity of a settlement, which is captured by MVC’s historic population density, is lost. That said, MVC’s core research question is mirrored in the current work: within a country, do people currently live and produce where they have historically lived and produced?

2 Data and Empirical Strategy

2.1 Data

Geolocalized Civilizations One of the primary contributions of the current work is the introduction of a new data set that geolocalizes all (or nearly all) recorded civilizations with a capitol and recording system.⁴ These locations and time frames come from a number of historic almanacs and encyclopedias; sources for which can be found in the appendix. The list of included civilizations is as comprehensive as possible as is the mapping of borders over time. To my knowledge, no other source attempts to locate all civilizations throughout history. The closest analogues are the *Centennia Historical Atlas*, which plots historic state boundaries for Europe, North African, and the Middle East starting in 1000 CE (Reed, 2016; Schönholzer and Weese, 2018); Murdock’s mapping of ethnic homelands (Michalopoulos and Papaioannou, 2013; Murdock, 1959); and the set of maps from Depetris-Chauvin (2015), who documents 72 civilizations across 15 time periods between 1000-1850 CE. In contrast, my data plot locations for roughly 1,470 civilizations over 500 periods from 3200 BCE to 2006 CE for all habitable continents.

These data are not perfect. Borders are likely mislabeled or effectively meaningless in many instances. That said, I believe this error to be unbiased (classical), leading to a simple attenuation in the coefficient of interest. It is highly unlikely that grid cells that were heavily populated historically as part of a civilization are mislabeled; instead, the converse is almost certainly true, especially given climatic change in the regions of the most historic civilizations—North Africa and the Middle East. Indeed, Classical Egypt predates

⁴The locations, sizes, placements, and sources of these borders were originally done by Andrew Tolleson. The maps were then used for an informational YouTube video: <https://www.youtube.com/watch?v=dp0tqdu7fH4>. These maps were then geolocalized by the SpARC lab at UC Merced, primarily by Erin Mutch and Amy Newsam.

the growth of the Saharan Desert into North Africa (Kröpelin et al., 2008). Furthermore, the early civilizations of the Middle East may have amplified desertification of this region, creating lightly populated/low output grids with a long state history (Ruddiman, 2003). In other words, bias from incorrectly measured civilization borders have likely *attenuated* the estimated relationship between state history and contemporary populations/lights and not led to a false-positive relationship.⁵

Sub-National State History The sub-national measure follows a similar form as measures in BCP, BOP, and Depetris-Chauvin. Using HSSW’s grid, I count the years a civilization has been within a particular cell. A simple intersection of a civilization with a cell is counted as the cell containing a civilization in that year. In other words, the civilization does not have to comprise a majority of area in a cell; if it is present, it is counted, regardless of coverage. This counting also excludes years when a particular grid is no longer part of a civilization, allowing for a more accurate counting of state history for marginal areas along the border or absences associated with a collapse of a civilization. Following BOP, I discount by 1% for 50 year intervals, sum the discounted years a cell has been part of a civilization, and set the variable relative to the maximum value.

Appendix Figure A.1 plots my measure of state history across the world. As expected, Eurasia has a longer history of a state presence; Africa shows a similar pattern to the previously collected data of Depetris-Chauvin with more state history in North Africa and the areas immediately beneath the Saharan desert; the New World shows patterns consistent with the big 3 civilizations—Mayan, Incan, and Aztec Empires; the United States shows a longer state history that is tied to its colonial origins; and Australia shows relatively recent state presence.

Outcome and Control Variables My hypothesis is that the historic location of civilizations have been sticky for contemporary development. To measure this association, I look at the relationship between sub-national state history and contemporary populations and output. To measure output, I follow HSSW and use their measure for the natural log of calibrated nighttime lights, but I also supplement HSSW’s DMSP nighttime light data with the generally improved VIIRS nighttime light data (Gibson, 2021). Population is from SEDAC’s UN adjusted population count data (CIESIN, 2018). HSSW’s DMSP nighttime light and SEDAC’s population are measured for 2010, while VIIRS nighttime lights are for 2014 (the first complete year of existence that corrects for stray lights; Gibson and

⁵Since my data are primarily based on historic atlases, civilizations are likely Euro-centric, affecting the included civilizations in both accuracy and the definition of a “civilization”. While not ideal, this problem is present in other databases like Murdock’s Ethnographic atlas.

Boe-Gibson (2021)).⁶

Control variables are also directly taken from HSSW, who use a set of 24 geoclimatic controls to account for the historic and persistent effect from agriculture and access to trade. The agricultural variables includes indicators for biome, mean temperature, mean precipitation, the number of grow days, land suitability, latitude, and elevation. Controls for trade include indicators for being on a coast, the mean distance to the coast, and indicators for being within 25km of a harbor, river, or lake. Summary statistics for all variables can be found in Table A.1.

I also use HSSW’s unit of analysis: quarter degree latitude by longitude cells for all habitable locations, totalling 242K observations. These grid cells, which are roughly 770 square kilometers at the equator (or $28\text{km} \times 28\text{km}$), are sufficiently large to overcome concerns of spatial spillovers while also providing a large number of observations per country, allowing for precise within country estimations.⁷

2.2 Validating Cell-Level State History

Figure 1 takes the average state history score within a country to compare my measure to the country-level measure of BCP and BOP. A clear and strong positive association exists between the two measures. The pairwise correlation is 0.77 and the Spearman rank correlation is 0.79. There are a few differences, however, between my measure of state history and the measure of BCP/BOP. First, BCP’s/BOP’s measure of state history codes for units of hierarchy beneath that of a kingdom or civilization. My measure only considers civilizations that have a capitol and a writing system and does not account for chiefdoms, or other lower-order forms of state organization, that are present in BCP’s/BOP’s data. Second, within country weighting for the geographic distribution of a historic civilization differs from BCP/BOP. Third, BCP/BOP discount invading or non-local civilizations, while I make no such distinction (the presence of a civilization is always coded as a “1” regardless of the civilization’s origins). To better highlight the differences between my measure and BCP/BOP’s measure, I examine two outliers: Ethiopia and Kyrgyzstan.

Case Analysis: Ethiopia and Kyrgyzstan Figure 2a plots the borders of the Kingdom of Axum (or Aksum) in 1 CE for modern day Ethiopia. As seen, the Kingdom of Axum covers a very small portion of Ethiopia, yet this kingdom is used to justify the long state history date of Ethiopia in BCP/BOP. Using the modern day borders of Ethiopia, most of

⁶VIIRS nighttime light data can be found at <https://eogdata.mines.edu/products/vnl/>. I use Annual VNL V2.1 data for the average masked score. The granular data are then summed within quarter degree grids as specified in Henderson et al. (2018).

⁷Spatial spillovers remain a concern; however, standard error adjustments provided in Table 1 suggest spillovers are not affecting the statistical significance of estimated effects of state history.

the country has had very little exposure to state organization, with only 3% being exposed to the Kingdom of Axum. This difference leads to Ethiopia being an outlier in Figure 2 and highlights the importance of a sub-national approach.

Figure b plots the modern day border of Kyrgyzstan as well as civilization borders for 1 CE. As seen a number of civilizations—the Wusun (NE), Han (SE), Dayuan (W), and Yuezhi (S) civilizations—overlap with Kyrgyzstan’s modern day borders at this early date, but BCP’/BOP’s measure ascribes a recent state history Kyrgyzstan. This is likely due to discounting from occupying instead of home-grown civilizations. For my purposes, this distinction is immaterial. I simply seek to measure the persistence of people and resulting cultures from exposure to higher order state organization.

2.3 Empirical Strategy

The empirical strategy follows that of BCP/BOP while using the unit of observation and geoclimatic controls of HSSW. This conditional correlation approach follows the literature and has similar threats to causation. The primary estimating equation is:

$$y_i = \alpha_i + \beta_{sh} \times StateHistory_i + \beta_2' \mathbf{Base}_i + \beta_3' \mathbf{Agr}_i + \beta_4' \mathbf{Trade}_i + \gamma_{ci} + \epsilon_i \quad (1)$$

Again, i represents land area divided into $0.25^\circ \times 0.25^\circ$ latitude by longitude grid. I consider three outcome variables, y_i : (1) the natural log of DMSP nighttime lights (from HSSW), (2) the natural log of population count (SEDAC; UN-WPP adjusted), and (3) VIIRS nighttime lights (c.2014).⁸ State history is my primary explanatory variable and measures the number of discounted years (relative to the maximum) a grid has been part of a civilization. The primary hypothesis is that populations and output have persisted in places that have historically been part of a civilization, implying $\beta_{sh} > 0$. The set of agricultural and access to trade variables from HSSW are respectively given by $Agr.$ and $Trade$. Country fixed effects are included in the baseline estimation and given above by γ_c .

Assuming agriculture is a necessary condition for state organization, the inclusion of HSSW’s agricultural set of controls will be collinear with state history. As noted in Table B.1, more favorable agricultural conditions are tied to state history, making it almost certain collinearity is affecting the estimated coefficient of state history. Nevertheless, favorable agricultural conditions likely also have an explanatory role for contemporary development outside of the relationship with historic development, and agricultural conditions alone do not fully explain the variation in state history. With this in mind, I include all of HSSW’s geoclimatic variables as controls but piecemeal introduce these controls to show

⁸The first full year of VIIRS availability is 2014. I use this year to stay as close as possible to other outcomes measured in 2010.

changes in the coefficient of state history. Following HSSW, I also report Shapley values in Appendix Table B.2, which decompose the marginal contribution to the R-square for each set of controls. These decompositions suggest state history independently explains a greater portion of the variation in nighttime lights and population than HSSW’s trade variables, but the set of agricultural variables remains the largest predictor of contemporary development. LASSO model selection includes all geoclimatic controls and state history, again suggesting an independent contribution from state history.

3 Results

3.1 Baseline Results

Base results are presented in Table 1. Panel A considers HSSW’s grid measure of nighttime lights; Panel B regresses the natural log of population from SEDAC (UN-WPP adjusted); and Panel C regresses the natural log of VIIRS nighttime lights. Controls are piecemeal introduced identically for all panels. Column (1) estimates the bivariate relationship between either the natural log of nighttime lights or population. Column (2) adds country fixed effects to the estimation of column (1). Column (3) introduces HSSW’s base variables—ruggedness and malaria ecology—to the estimation of column (2). Columns (4) and (5) separately include the set of agriculture and access to trade controls to the estimation of column (3). And column (6) comprises the base specification, which includes all of HSSW’s geoclimatic controls and country fixed effects. Three standard errors are included: (1) country clustered, (2) three-by-three grid clustered (from HSSW), and (3) spatially adjusted for 200 km. I use the highest nested (i.e., country) and most conservative cluster for determining statistical significance. The coefficient of state history is statistically significant at the 1% level for all specifications and all outcomes.

For the bivariate estimates of column (1) with no country fixed-effects, a standard deviation increase in state history (i.e., 0.2) is associated with roughly a tripling of DMSP nighttime lights, a five-fold increase in population, and a tripling of VIIRS nighttime lights. These large effects are in part due to the large number of cells with no nighttime lights; an issue that is discussed in Sections 3.3. The estimated effect of state history is effectively unaltered when including country fixed effects in column (2), a point also discussed in greater detail below. The inclusion of HSSW’s base controls—i.e., ruggedness and malaria ecology—in column (3) again do not significantly alter the coefficient of state history.

The inclusion of agricultural controls in column (4), however, causes the coefficient of state history to be reduced by roughly half for all outcomes, implying a one standard deviation increase in state history is now associated with a 100% increase in DMSP nighttime lights, a 140% increase in population, and a 110% increase in VIIRS nighttime lights. As discussed

in Section 2.3, HSSW’s agricultural controls intend to capture agricultural conditions that have led to prolonged state history, suggesting the inclusion of the agricultural set of controls splits a common effect among many coefficients. That said, the lowered magnitude coefficient of state history from the inclusion of agricultural controls remains statistically significant and the lowered magnitude still suggest a sizable economic impact from a grid having more state history. Furthermore, Appendix B provides additional evidence that state history is indeed making an independent and sizable contribution in explaining the three outcomes of interest.

Column (5) adds HSSW’s trade variables to the regression specification of column (3)—i.e., country FE and the base set of controls—with no substantial changes to the coefficient of state history, and column (6) comprises the baseline specification and includes all controls. The estimated coefficient of column (6) shows the reduced magnitude associated with the introduction of agricultural variables in column (4) with point estimates and economic effects being similar in magnitude to those in column (4).

The relatively large effects of state history are inline with the estimated effects of other variables. For example, from Tables A.1 and column (6) of Tables E.1 and E.2, a one standard deviation increase in the average monthly temperature is associated with roughly a quadrupling of DMSP nighttime lights and a quintupling of population. Similarly large effects are estimated for other covariates.

3.2 Institutions or Agglomeration

No Alteration from Country FE Table 1 provides the coefficients of state history for identical estimations absent country fixed effects. The magnitudes are surprisingly similar, and a Hausman-based test shows no statistically significant differences in the coefficients of state history from a model with country fixed effects and a model omitting country fixed effects for both nighttime light outcomes; although, statistically significant differences do exist (larger when omitting country FE) when considering population and including agricultural controls (columns 4 and 6). The consistency in magnitude across specifications suggests time-invariant country-level factors like institutional quality are not playing a large role in the relationship between a cell’s state history and its corresponding lights/population. In other words, if state history were working solely through country-level institutional quality, one would expect country-level fixed effects to alter (likely attenuate) the relationship between state history and nighttime lights/population. The absence of any change from the inclusion of country fixed effects suggests any country-level differences are orthogonal to the estimated grid-cell relationship.

Institutional Quality FE While the the estimated relationship between cell-level state history and nighttime lights/population appears to be independent of country-level differences, country FE may possibly over control for other factors not directly related to institutional quality. With this idea in mind, Table 2 directly controls for institutional quality using data from the Polity5 Project (Marshall and Gurr, 2020). The Polity5 Project documents country-level institutional changes from 1800 CE to the present and categorizes institutions based on autocratic (-10; “strongly autocratic”) versus democratic (+10; “strongly democratic”) governance in the “POLITY” variable. In addition to this 21-point scale, countries are take differing values based on transitions from one type of governance to another. Table 2 uses the POLITY data for 2010 as a direct control for institutional quality and performs heterogeneity tests for a varied effect of state history by institutional quality.

The autocracy/democracy variation is closely tied to the arguments of BOP. The idea being that older lived autocratic governments have persisted, serving as a (negative) mechanism connecting state history to current economic development (Hariri, 2012; Olsson and Paik, 2020). Alternatively, newer states developed inclusive, democratic institutions that led to economic prosperity; BOP’s positive mechanism of state history. With these mechanisms in mind, the POLITY variable is a good candidate to measure the institutional quality mechanisms proposed by BOP.

Column (1) of Table 2 directly controls for the POLITY score by introducing each category as a fixed effect.⁹ The inclusion of institutional quality fixed effects does not significantly alter the coefficient of state history for any of the three outcomes; the estimated effects are similar in magnitude to those of column (6) of Table 1, with or without country FE.

Columns (2)-(4) use a dichotomized POLITY score that serves as a democracy indicator—i.e., $POLITY > 0$. Countries with transition POLITY scores are omitted. Column (2) directly controls for the country-level indicator of democracy, leading to no large differences in the coefficients of state history. Column (3) interacts the democracy indicator with state history to test whether the effect of state history differs within democratic versus autocratic governments. For all panels, the coefficient of the interaction is insignificant, suggesting there is no statistically significant difference in the effect of state history across democratic or autocratic countries. Column (4) builds on the estimates of column (3) by including country fixed effects.¹⁰ As with column (3), however, no statistically significant difference is seen in the effect of state history for democratic countries.

⁹Linearly controlling for POLITY does not alter the estimated coefficient of state history.

¹⁰Note that the democracy indicator is no longer included due to being perfectly collinear with the country FE, but the interaction is estimated.

Relationship Driven by Cities Tables 1 and 2 provide evidence that the relationship between state history and nighttime lights/population is unaffected by country level controls, suggesting a federal institution mechanism is absent. Instead of “proving a negative”, Table 3 provides evidence for agglomeration by showing the effect of state history is driven by proximity to urban areas.¹¹

Table 3 uses the mid-point of a cell to determine the distance (in 100s of kilometers) to the closest urban location. This distance is then interacted with the cell’s state history score to determine how the marginal effect of state history on each of the three outcomes changes with distance from a city. Column (1) includes all baseline controls; column (2) adds country FE to the specification of column (1); and column (3) adds closest city FE to the specification of column (2). In general, the estimates of Table 3 show that the effect of state history is significantly tied to urban distance, becoming weaker in more rural areas and equal to zero roughly 500-600km away from a city. The decaying effect of state history holds for all specifications, and the inflection point, or where the marginal effect of state history is zero, is relatively consistent.

Agglomeration is much more pronounced in cities (Chanda and Ruan, 2017; Glaeser, 2010), and proximity to a city is driving the positive and significant effect of state history. This is counter to the estimates of Table 2, which show no variation in the effect of state history by federal-level institutions. The estimates from the heterogeneity tests of Table 2 and 3 and the lack of impact from country FE in Table 1 suggest a stronger role for an agglomeration mechanism than an institutional quality effect.

3.3 Robustness and Alternative Specifications

Unlit/Unpopulated Cells Figures C.1 - C.4 respectively plot the distributions for HSSW’s DMSP nighttime lights, SEDAC’s population count, VIIRS nighttime lights, and state history. Roughly 60% of cells are unlit (both DMSP and VIIRS) and 10% of cells are unpopulated. The sizable frequency of “empty” cells is likely to affect estimation, so Table C.1 considers alternative samples that exclude unlit/unpopulated cells, which comprises those in the bottom decile, and those in the top decile.¹²

Column (1) of Table C.1 replicates the estimation of column (6) of Table 1 while omitting the bottom decile of each outcome measure, which comprises grid cells with no lights (Panels A and C) or no population (Panel B). For nighttime lights in Panel A and C, this leads to a large reduction in the sample size (again, 60% of grids have no lights) and a reduction in

¹¹Urban location data are from Reba, Reitsma, and Seto (2016), which aggregates data from Chandler (1987) and Modelski (1999).

¹²Bias from the large number of unlit/unpopulated cells is likely to be positive. Considering a positive linear fit, the large number of zero, or low value, observations will increase weight in the lower, lefthand side, potentially creating upward bias in the estimated relationship of interest.

the coefficient of state history that remains positive and statistically significant at the 1% level. For population in Panel B, the omission of unpopulated grid cells has a minor (to nonexistent) effect on the magnitude and significance on the coefficient of interest.

Column (2) of Table C.1 omits instead the top decile of nighttime lights and population. The omission of the top decile effectively removes urban areas and shows that my sub-national measure of state history is not only accounting for the persistence of cities. The argument being that the persistence of both urban areas and areas in the hinterlands that feed these cities and more rural areas that are taxed and have access to public goods like roads to cities are all accounted for within a civilization’s boundary. The hypothesis is not tied strictly to urban areas; although, effects are more pronounced in these areas due to agglomeration. As expected, estimates from column (3) are smaller in magnitude compared to the baseline estimates of column (6) in Table 1 but remain positive and statistically significant, suggesting contemporary urban areas are not the sole cause of the positive relationship between state history and contemporary output/populations.¹³

Column (3) omits both the top and bottom quartile of contemporary nighttime lights and population. For both nighttime light measures (Panels A and C), the coefficient of state history is reduced—primarily from unlit cells—but remains significant at the 1% level; Panel B shows a relatively smaller attenuation when considering population, which is driven by the omission of more populated cells.

To summarize, prevalent unlit/unpopulated cells do not dissipate the positive and statistically significant association between output/population and state history, but the sheer volume of these cells does significantly affect the magnitude of the relationship for nighttime lights. From columns (2) and (3) the effect of state history is not driven solely by urban areas, suggesting a rural/hinterland persistence that has not been documented in other work.

Large Country Effects While country FE account for country-specific time-invariant factors, the quarter degree latitude/longitude grids are mechanically tied to the area of a country. This implies larger countries will have more observations and a greater weighting on the estimated coefficient, with or without country FE. To show that this unequal weighting is not driving the relationship between state history and nighttime lights/population, Table D.1 limits the sample to within country observations for the largest countries—Russia, Canada, China, and the United States—and omits these large countries from the base sample.

Columns (1)-(4) respectively limit the sample to cells within Russia, Canada, China, and the United States. Point estimates vary from the baseline estimates of Table 1, but

¹³Similar effects are observed when omitting city locations from Table 3

this is expected given the focus on singular countries; however, the estimated coefficients from omitting these large countries in column (5) are similar in magnitude to the baseline estimates, suggesting little influence from a singular, but heavily weighted, country.

Effect by Continent Table D.2 examines the baseline estimation of Table 1, column (6) by continent. For nighttime lights, a positive and statistically significant effect of state history is estimated for Africa, Asia, South America, and Oceania. The largest estimated coefficient is for South America—inline with Maloney and Valencia Caicedo (2014), but this large coefficient is in part due to the lower standard deviation for state history within South America (see Table A.1). That is, a standard deviation increase in state history for South America is associated with roughly a doubling in nighttime lights, a finding very similar to the baseline estimate of column (6) in Table 1. Estimated effects for population are very similar in Panel B. To account for differences in the variation of state history across continents, Table ?? standardizes state history within continent (or Old/New World grouping) to a mean of zero and standard deviation of one. Once this adjustment is made, the effect of state history is mostly inline across continents, still remaining relatively large in South America and the Old World effect of state history is more pronounced when compared to the New World effect.

4 Conclusion

This paper introduces a novel data set that documents the location and time of a civilization’s borders. These borders are not tied to current country borders and allow for the creation of a measure of state history by counting the years a plot of land has been exposed to any civilization. Plots are from a 0.25° by 0.25° fishnet created by Henderson et al. (2018) that provides a large number of observations per country. I then test the relationship between nighttime lights and population both across and within countries, finding a consistent (between and within country), statistically significant, and positive role for state history in measuring differences in economic output today.

The consistency in the effect of state history with and without country fixed effects suggests country-level factors, like institutional quality, are not influencing the estimated relationship. This contrasts with the mechanism posited by BCP and BOP that suggests improvements in state capacity over time lead to better economic outcomes today. Furthermore, directly controlling for institutional quality or testing for heterogeneity from institutional quality shows no role for federal level institutions. Instead, it is much more likely that the mechanism connecting state history to contemporary development is agglomeration. Indeed, I observe statistically significant heterogeneity in the effect of state

history tied to the distance from an urban center, where agglomeration effects are much more pronounced. The lack of heterogeneity from institutional quality measures coupled with the strong heterogeneity from urban distances suggests agglomeration as the much more likely mechanism. That said, I cannot explicitly rule out federal institutional quality as a primary mechanism.

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5 Tables and Figures

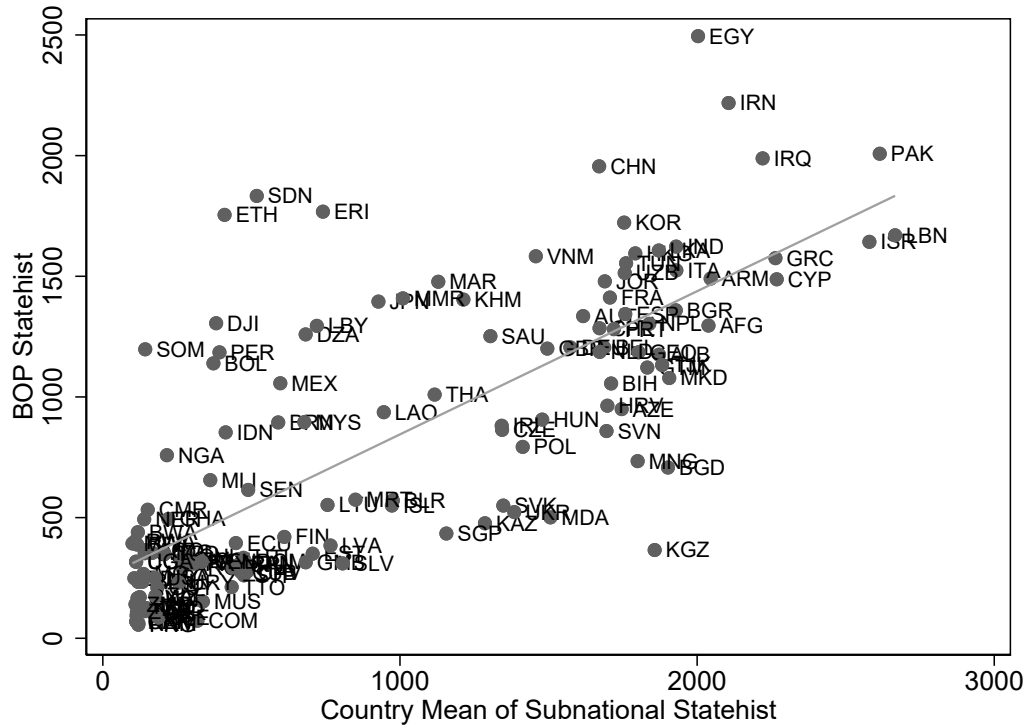
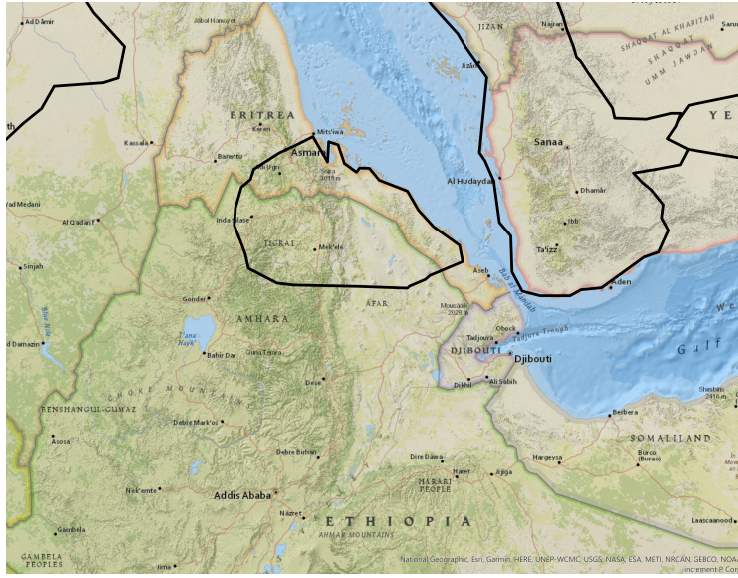


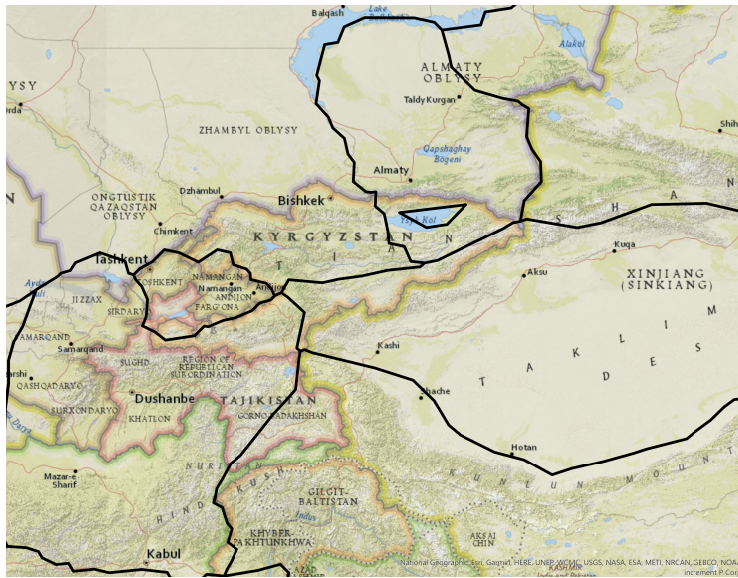
Figure 1: Comparing BOP’s State History to Cell State History

Notes:

This figure plots state history from Borcan et al. (2018; y-axis) relative to a country aggregated (mean) measure of my cell-level state history measure based geolocalized civilizations (x-axis). The two measures are highly correlated with a pairwise correlation coefficient of 0.77 ($p < 0.000$). Of note, outliers from Figure 1 illustrate the differences between my measure and that of BOP. This is discussed further in the case analysis of Section 2.1 and in Figure 2.



(a) Ethiopia



(b) Kyrgyzstan

Figure 2: Civilization Borders: 1 CE

Notes:

This figure plots geolocalized civilization borders for 1 CE relative to the modern day borders of Ethiopia (sub-figure (a)) and Kyrgyzstan (sub-figure (b)). These two countries serve as outliers in Figure 1, showing large differences between my measure of state history and that of BOP. Ethiopia has a long state history in BOP but relatively short state history in my analysis. This disparity is due to the geographic distribution of the Kingdom of Aksum, the borders of which for 1 CE are plotted on top of modern day Ethiopia in sub-figure (a). As seen, the Kingdom of Aksum only covers a small portion of modern day Ethiopia, with the vast majority of the country not belonging to a historic civilization. At the opposite end of the spectrum, BOP define Kyrgyzstan as having a relatively recent state history, while my measure shows a longer lived state presence. Sub-figure (b) plots civilizations from 1 CE around modern day Kyrgyzstan. While being in modern-day Kyrgyzstan, these civilizations are not local to Kyrgyzstan, leading to substantial discounting in BOP's measure of state history.

Table 1: Baseline Results: Nighttime lights, population, and state history

	(1)	(2)	(3)	(4)	(5)	(6)
Panel A. ln DMSP Nighttime Lights, 2010						
State History	5.7222***	8.0480***	8.0173***	3.4938***	7.9298***	3.5624***
country-clustered s.e.	(0.9966)	(0.8897)	(0.8573)	(1.0009)	(0.9548)	(0.8074)
HSSW (3-grid) clustered s.e.	(0.0515)	(0.0989)	(0.0961)	(0.0730)	(0.0928)	(0.0705)
spatial-200km-clustered s.e.	(0.3559)	(0.6368)	(0.6121)	(0.3862)	(0.5881)	(0.3569)
Observations	242,184	242,184	242,184	242,184	242,184	242,184
R Sqr.	0.1408	0.3965	0.4068	0.5747	0.4196	0.5857
β^{SH} , omitting country FE		5.7222***	5.7001***	2.9039***	6.2498***	3.1713***
		(0.9966)	(1.1000)	(0.8411)	(1.0816)	(0.8396)
p-value of difference		0.0898	0.0920	0.3776	0.2015	0.5204
Panel B. ln Population, 2010						
State History	8.4013***	9.4568***	9.4860***	4.2796***	9.4541***	4.4186***
country-clustered s.e.	(1.8863)	(1.4901)	(1.4450)	(0.5559)	(1.5534)	(0.4679)
HSSW (3-grid) clustered s.e.	(0.0917)	(0.1662)	(0.1610)	(0.1111)	(0.1604)	(0.1090)
spatial-200km-clustered s.e.	(0.4546)	(0.6929)	(0.6702)	(0.3812)	(0.6588)	(0.3670)
Observations	242,184	242,184	242,184	242,184	242,184	242,184
R Sqr.	0.2100	0.6147	0.6168	0.7749	0.6238	0.7809
β^{SH} , omitting country FE		8.4013***	9.7261***	6.7255***	9.9031***	6.8219***
		(1.8863)	(1.7672)	(0.9417)	(1.6190)	(0.9180)
p-value of difference		0.5940	0.8986	0.0210	0.8086	0.0135
Panel C. ln VIIRS Nighttime Lights, 2014						
State History	5.7467***	7.9630***	7.9365***	3.7159***	7.9128***	3.8184***
country-clustered s.e.	(1.0050)	(0.9092)	(0.8898)	(1.0285)	(0.9718)	(0.8424)
HSSW (3-grid) clustered s.e.	(0.0773)	(0.1605)	(0.1584)	(0.1229)	(0.1555)	(0.1186)
spatial-200km-clustered s.e.	(0.3203)	(0.6109)	(0.5966)	(0.3929)	(0.5737)	(0.3608)
Observations	242,184	242,184	242,184	242,184	242,184	242,184
R Sqr.	0.1471	0.3843	0.3895	0.5426	0.3977	0.5499
β^{SH} , omitting country FE		5.7467***	5.6908***	2.9386***	6.1925***	3.2094***
		(1.0050)	(1.1090)	(0.8713)	(1.1117)	(0.8692)
p-value of difference		0.1072	0.1080	0.4583	0.2029	0.5138
Controls (for all panels):						
Country FE		Y	Y	Y	Y	Y
Base variables			Y	Y	Y	Y
Agricultural variables				Y		Y
Trade variables					Y	Y

Summary & Notes: This table presents the baseline estimates showing that sub-national state history has a persistent and positive relationship with contemporary output (proxied by nighttime lights) and population. Panel A regresses the natural log of HSSW's DMSP nighttime lights; Panel B regresses the natural log of cell-level population; and Panel C regresses the log of VIIRS nighttime lights. Sets of controls and summary statistics are given in Table A.1. Statistical significance is determined by country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *. Spatially adjusted (200km) and HSSW's 3-grid cluster standard errors are also reported.

Table 2: Institutional Quality: Controlling and Testing Heterogeneity

	(1)	(2)	(3)	(4)
Panel A. ln DMSP Nighttime Lights, 2010				
State History	3.533*** (0.721)	3.592*** (0.800)	3.623*** (0.625)	3.250*** (0.590)
Polity Indicator, 1[Democracy]		0.577** (0.236)	0.590* (0.352)	
State History × Polity Ind.			-0.045 (0.892)	0.448 (1.423)
Observations	238,952	236,120	236,120	236,119
R Sqr.	0.512	0.500	0.500	0.584
Panel B. ln Population, 2010				
State History	5.113*** (0.602)	6.606*** (0.839)	3.792*** (0.772)	3.997*** (0.643)
Polity Ind., 1[Democracy]		-0.336 (0.384)	-1.528*** (0.575)	
State History × Polity Ind.			4.127*** (1.302)	0.862 (0.953)
Observations	238,952	236,120	236,120	236,119
R Sqr.	0.707	0.647	0.656	0.775
Panel C. ln VIIRS Nighttime Lights, 2014				
State History	3.624*** (0.725)	3.760*** (0.828)	4.031*** (0.628)	3.534*** (0.748)
Polity Ind., 1[Democracy]		0.692*** (0.233)	0.807** (0.361)	
State History × Polity Ind.			-0.397 (0.928)	0.476 (1.595)
Observations	238,952	236,120	236,120	236,119
R Sqr.	0.487	0.476	0.476	0.549
Controls (for all panels):				
Polity FE	Y			
Country FE				Y
All geoclimatic controls	Y	Y	Y	Y

Summary & Notes: This table tests whether direct controls for country-level institutional quality affect the estimated coefficient of state history across the two nighttime lights measures (Panels A and C) and population (Panel B). As shown, controlling for institutional quality (done by categorical fixed effects in column 1) does not alter the estimated effect. Going further, columns (2)-(4) perform heterogeneity tests using an indicator of democracy constructed from the Polity variable used in column (1). Estimates suggest the effect of state history is not significantly affected by the type of country-level governance. This suggests a channel outside of state capacity is driving the baseline estimates of Table 1. Geoclimatic controls are from HSSW and listed in Table 1. Statistical significance is determined by ²¹country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *.

Table 3: Decay from Urban Distance

	(1)	(2)	(3)
Panel A. ln DMSP Nighttime Lights, 2010			
State History	3.3166*** (0.9302)	4.2156*** (0.8080)	3.7663*** (0.7025)
Distance to nearest city (in 100km)	-0.0863*** (0.0256)	-0.0271 (0.0215)	-0.0243 (0.0157)
State History \times Distance	-0.5934*** (0.1733)	-0.7732*** (0.1874)	-0.6755*** (0.1717)
Inflection point	559km	545km	558km
Observations	242,184	242,180	242,153
R Sqr.	0.5209	0.5969	0.6551
Panel B. ln Population, 2010			
State History	5.0229*** (0.7167)	4.5366*** (0.7354)	4.3271*** (0.4674)
Distance to nearest city (in 100km)	-0.1817*** (0.0403)	-0.0759*** (0.0218)	-0.0560*** (0.0179)
State History \times Distance	0.4185 (0.4467)	-0.5692* (0.3432)	-0.6357** (0.2509)
Inflection point	n.a.	797km	681km
Observations	242,184	242,180	242,153
R Sqr.	0.6786	0.7911	0.8375
Panel C. ln VIIRS Nighttime Lights, 2014			
State History	3.5369*** (0.9733)	4.6404*** (0.8471)	4.0630*** (0.6619)
Distance to nearest city (in 100km)	-0.0708*** (0.0241)	-0.0125 (0.0195)	-0.0178 (0.0144)
State History \times Distance	-0.6168*** (0.1803)	-0.7827*** (0.1958)	-0.7358*** (0.1783)
Inflection point	573km	593km	552km
Observations	242,184	242,180	242,153
R Sqr.	0.4911	0.5603	0.6240
Controls:			
Country FE		Y	Y
City FE			Y
All geoclimatic controls	Y	Y	Y

Summary & Notes: This table provides evidence the persistent effects of state history are largely due to urban proximity. The coefficient of the interaction shows the marginal effect of state history decays as distance from a current or historic city increases; the effect of state history is indistinguishable from zero at roughly 600 km from an urban location. Panel A regresses the natural log of HSSW's measure of nighttime lights based on DMSP satellite readings for 2010; Panel B regresses the natural log of a cell's total population in 2010 (SEDAC); and Panel C regresses the natural log of VIIRS nighttime light for 2014, the first complete year of data with stray light corrections. Geoclimatic controls are from HSSW and listed in Table 1. Statistical significance is determined by country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *.

Online Appendix Tables and Figures

A Summary Statistics and Additional Figures

Table A.1: Summary Statistics

Variable	Obs.	Mean	Std Deviation	Min	Max
State History	242,184	0.1964	0.2045	0	1
Years, Discounted (1% per 50)	242,184	665.1138	692.5049	0	3386.872
Years, Not Discounted	242,184	784.6688	900.0418	0	5206
By continent:					
Africa	41,346	0.1178	0.1723	0	1
Asia	51,258	0.4633	0.1806	0	0.9737
Europe	60,050	0.2252	0.1509	0	0.7704
South America	25,369	0.0860	0.0342	0.0013	0.2905
North America	50,941	0.0533	0.0622	0	0.6966
Oceania	13,220	0.0400	0.0051	0.0179	0.0630
Outcome Variables:					
HSSW's ln DMSP Nighttime Lights (c.2010)	242,184	-3.3571	3.1186	-5.6841	6.9410
SEDAC's ln Population Count (c.2010)	242,184	6.2954	3.7458	0	16.1906
ln VIIRS Nighttime Lights (c.2014)	242,184	2.3087	3.0636	0	14.6335
Control Variables:					
Base set:					
Rugged	242,184	2.7806	4.8515	0	95.8144
Malaria Ecology	242,184	1.9208	5.2886	0	38.0810
Agriculture set:					
Biome 1 (tropical moist forest)	242,184	0.1168	0.3211	0	1
Biome 2-3 (tropical dry forest)	242,184	0.0223	0.1477	0	1
Biome 4 (temperate broadleaf)	242,184	0.1044	0.3059	0	1
Biome 5 (temperate conifer)	242,184	0.0330	0.1785	0	1
Biome 6 (boreal forest)	242,184	0.1664	0.3724	0	1
Biome 7-9 (tropical grassland)	242,184	0.1208	0.3259	0	1
Biome 8 (temperate grassland)	242,184	0.0772	0.2670	0	1
Biome 10 (montane grassland)	242,184	0.0334	0.1797	0	1
Biome 11 (tundra)	242,184	0.1221	0.3274	0	1
Biome 12 (Mediterranean forest)	242,184	0.0242	0.1536	0	1
Biome 13 (desert)	242,184	0.1753	0.3802	0	1
Biome 14 (mangroves)	242,184	0.004	0.0634	0	1
Temperature	242,184	10.0181	13.7679	-22.2861	30.3660
Precipitation	242,184	60.8169	59.2736	0.3866	921.9088
Growing Days	242,184	139.63	99.0428	0	366
Land Suitability	242,184	0.2748	0.3200	0	1
Absolute Latitude	242,184	38.3148	20.9350	0.125	74.875
Elevation	242,184	0.6045	0.7899	-0.1873	6.1690
Trade set:					
Coastal Ind.	242,184	0.0972	0.2963	0	1
Dist. to Coast	242,184	0.4865	0.4811	0	2.2738
Harbor w/in 25km	242,184	0.0273	0.1629	0	1
River w/in 25km	242,184	0.0273	0.1629	0	1
Lake w/in 25km	242,184	0.0108	0.1035	0	1

Notes:

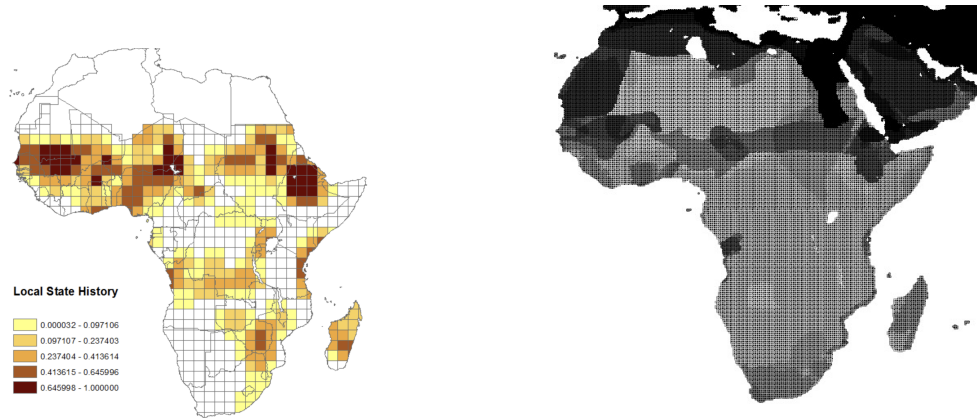
This table provides summary statistics. The unit of observation for all variables is the 0.25° by 0.25° grid cell from Henderson et al. (2018).



Figure A.1: Density Plot (Heatmap) of State History

Notes:

This figure plots the cell-level gradient of my state history measure. Darker areas represent longer "civilized" grids.



(a) Depetris-Chauvin (Figure 4; 2015)

(b) Civ Based

Figure A.2: Comparing State History Measures for Sub-Saharan Africa

Notes:

Sub-figure (a) plots Depetris-Chauvin's constructed measure of state history (Figure 4 from Depetris-Chauvin (2015)), while sub-figure (b) focuses on Africa from Figure A.1. The measures show considerable overlap.

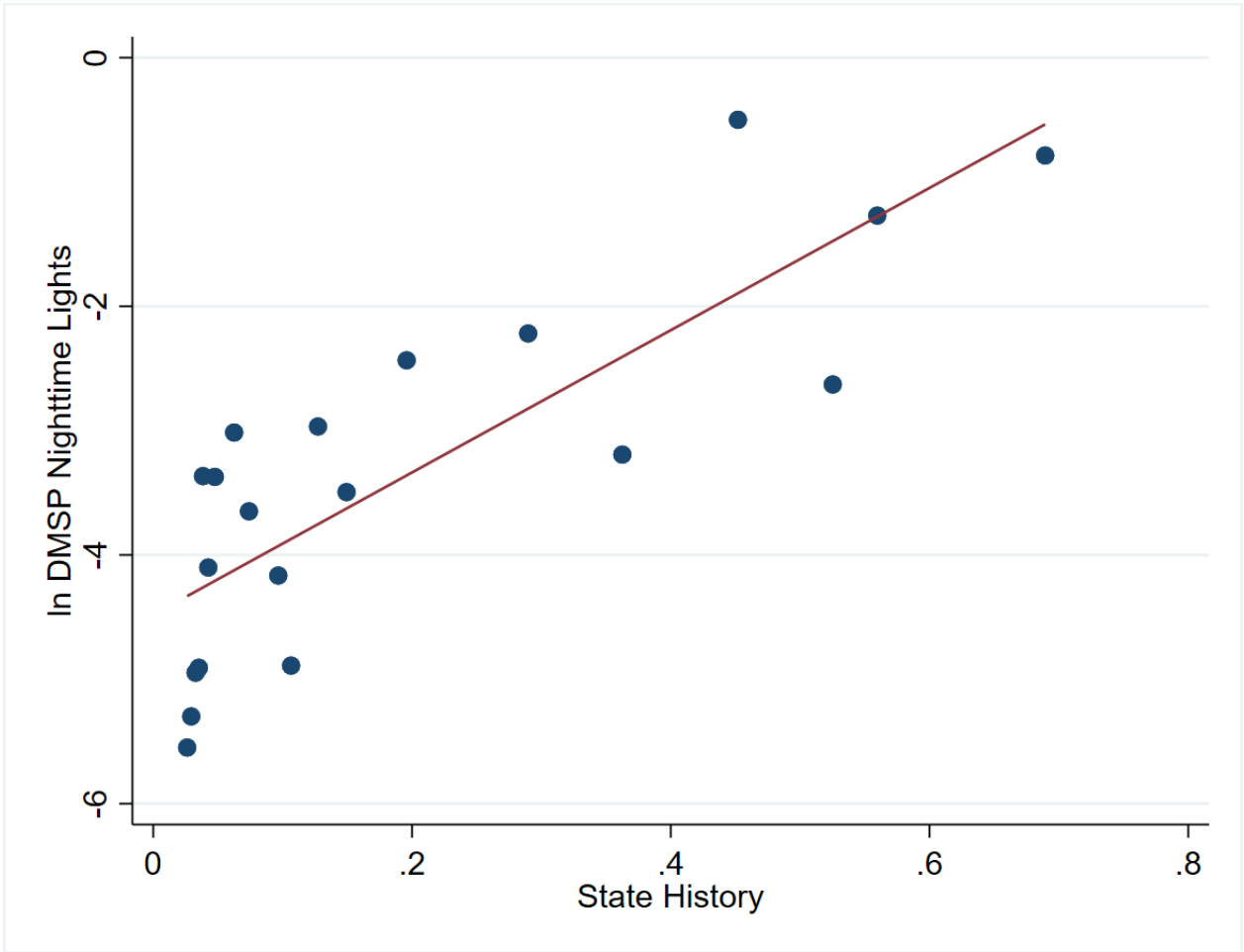


Figure A.3: Binscatter: State History and HSSW's ln DMSP Nighttime Lights

Notes:

This figure gives the cell-level bivariate binned scatter plot between the natural log of DMSP nighttime lights and state history.

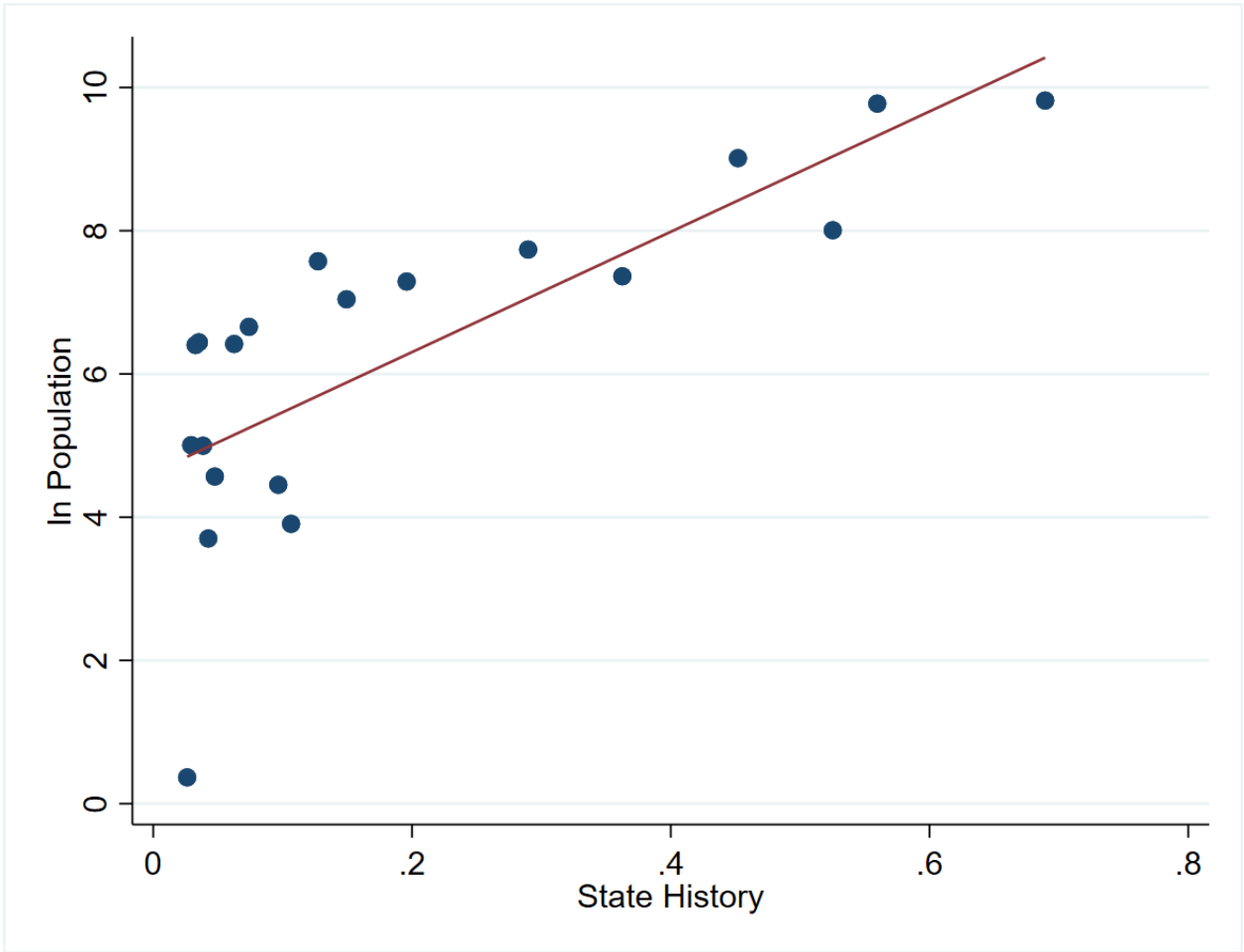


Figure A.4: Binscatter: State History and ln Population

Notes:

This figure gives the cell-level bivariate binned scatter plot between the natural log of SEDAC's population and state history.

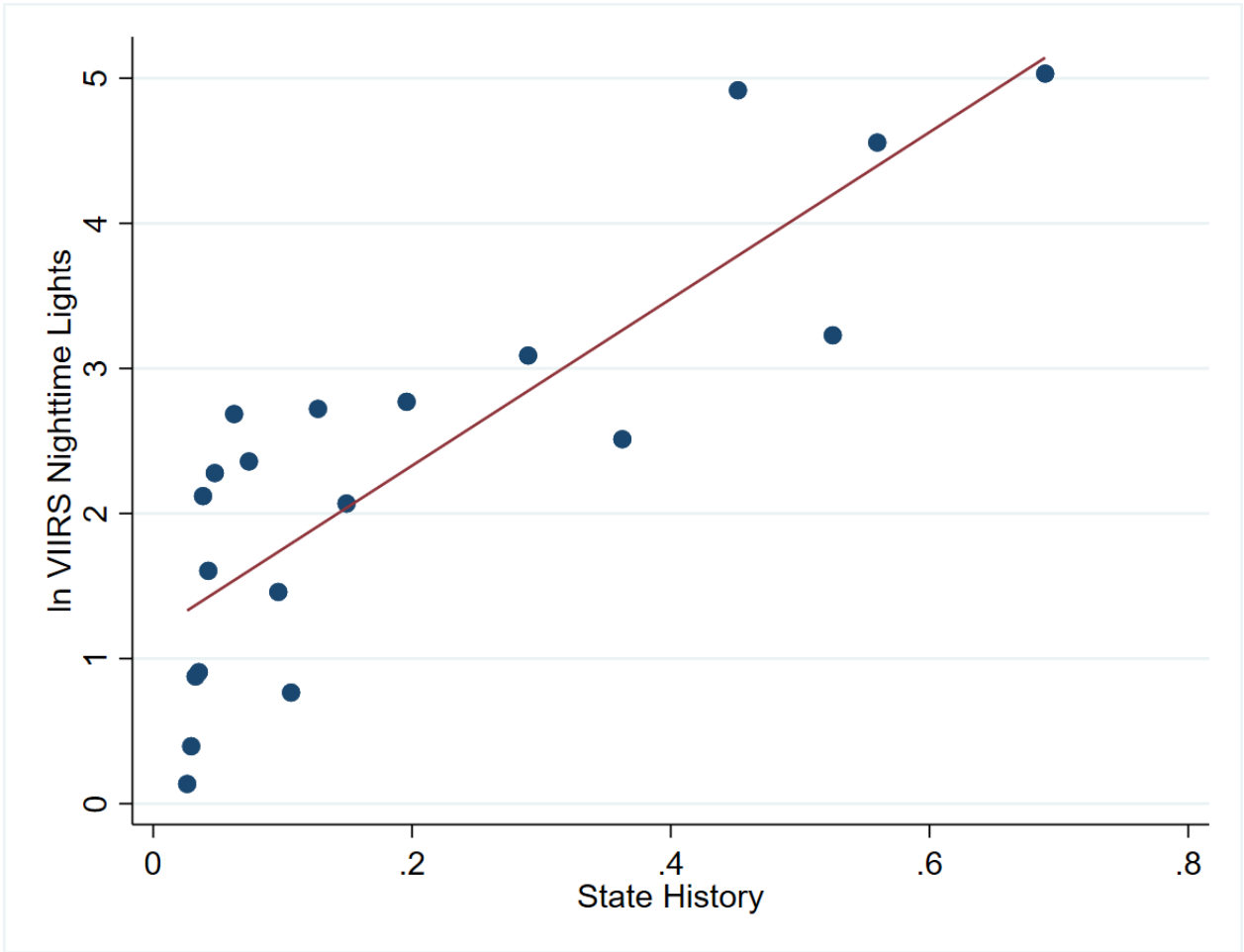


Figure A.5: Binscatter: State History and ln VIIRS Nighttime Lights

Notes:

This figure gives the cell-level bivariate binned scatter plot between the natural log of VIIRS nighttime lights and state history.

B Agriculture and the Independent Contribution of State History

Table B.1: Partial Correlation Matrix (State History Only)

Variable	Pairwise Correlation Coefficient with State History
Base set:	
Rugged	0.1594
Malaria Ecology	-0.2062
Agriculture set:	
Biome 1 (tropical moist forest)	-0.0450
Biome 2-3 (tropical dry forest)	0.0745
Biome 4 (temperate broadleaf)	0.2101
Biome 5 (temperate conifer)	0.0224
Biome 6 (boreal forest)	-0.1492
Biome 7-9 (tropical grassland)	-0.2175
Biome 8 (temperate grassland)	0.0502
Biome 10 (montane grassland)	0.1289
Biome 11 (tundra)	-0.2211
Biome 12 (Mediterranean forest)	0.1468
Biome 13 (desert)	0.1974
Biome 14 (mangroves)	-0.0022
Temperature	0.0860
Precipitation	-0.1186
Growing Days	-0.0579
Land Suitability	0.2040
Absolute Latitude	-0.0029
Elevation	0.1687
Base set:	
Coastal Ind.	-0.0433
Dist. to Coast	0.1917
Harbor w/in 25km	0.0308
River w/in 25km	0.0370
Lake w/in 25km	-0.0161

Summary & Notes: This table provides correlation coefficients for all control variables and state history. All pairwise correlations are statistically significant ($p < 0.01$) except for absolute latitude ($p = 0.16$) and the Mangrove Biome Indicator ($p = 0.27$).

Table B.2: Shapley Values

Outcome:	ln DMSP Nighttime Lights		SEDAC's ln Population Count		ln VIIRS Nighttime Lights	
	Shapley Value	Percent	Shapley Value	Percent	Shapley Value	Percent
State History	0.06	10.83%	0.10	12.90%	0.07	12.06%
Base	0.01	1.45%	0.03	3.61%	0.01	1.22%
Agriculture	0.29	49.02%	0.32	41.54%	0.27	49.02%
Trade	0.03	4.62%	0.01	1.18%	0.02	3.17%
Country FE	0.20	34.09%	0.32	40.77%	0.19	34.53%
Total	0.5855	100%	0.7909	100%	0.5500	100%

Summary & Notes: This table provides Shapley values as in Henderson et al. (2018). State history is shown to independently account for roughly 10% of the explained variation (i.e., R^2) for each of the three outcomes; this is behind only the set of country FE and the set of agricultural controls.

Table B.3: Lasso Control Ranking:
Top 5 Penalized Coefficients (Standardized; Rank in Parantheses)

Outcome:	ln DMSP Nighttime Lights	ln Population	ln VIIRS Nighttime Lights
State History	0.73 (4)	0.91 (4)	0.06 (2)
Temperature	1.29 (1)	1.62 (1)	
Australia Indicator	-0.81 (2)	-0.98 (2)	-0.04 (4)
Number of Growing Days	0.79 (3)	0.97 (3)	0.08 (1)
Land Suitability	0.69 (5)	0.65 (5)	
Precipitation			-0.05 (3)
Abs. Latitude			-0.04 (5)

Summary & Notes: This table uses a LASSO cross-validation exercise to estimate penalized coefficients for all standardized variables (i.e., mean=0, s.d.=1). The penalized coefficient for (standardized) state history ranges from 2nd to 4th in magnitude out of the 204 variables considered. Additional tests for a LASSO selection of coefficients (double-selection and cross-fit partialing-out) list all used variables as relevant for estimation. This indicates that the state history variable provides valuable variation independent from the set of agricultural controls.

C Robustness to Unlit/Unpopulated Cells

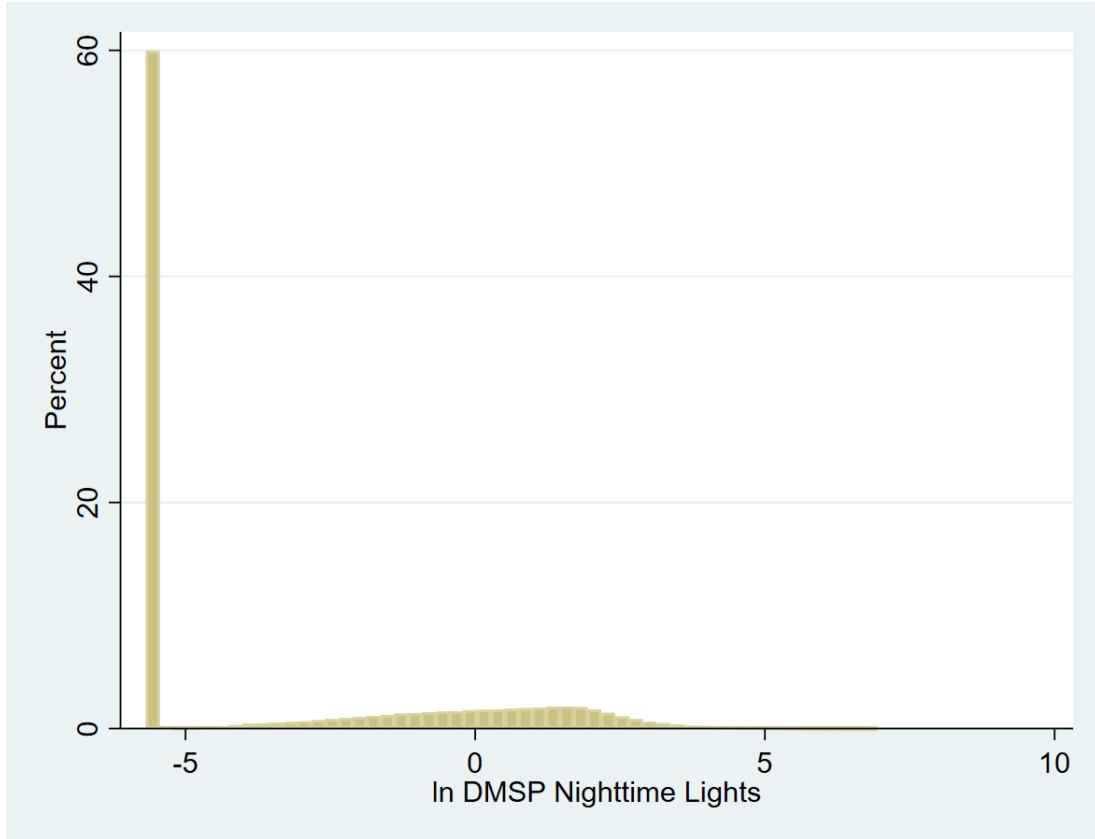


Figure C.1: Histogram of Grid DMSP Nighttime Lights

Notes:

This figure shows the distribution of DMSP nighttime lights from Henderson et al. (2018). Notice that roughly 60% of values are bottom coded.

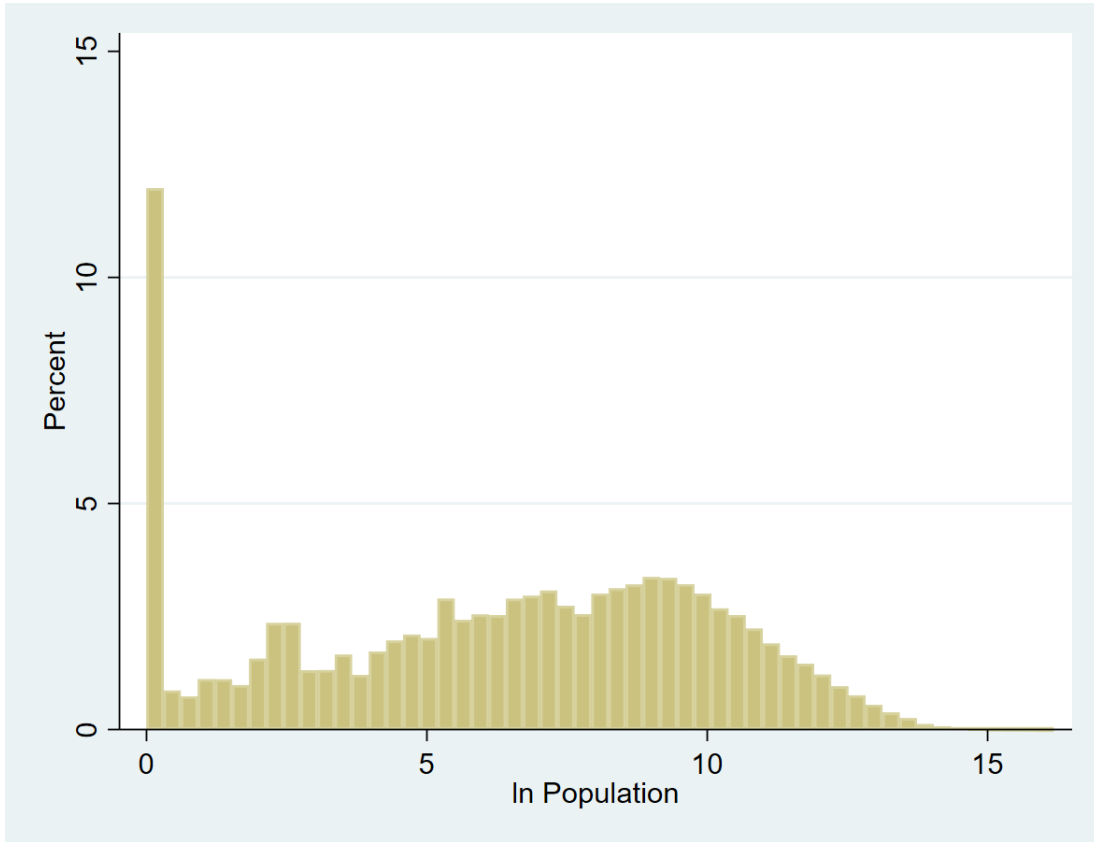


Figure C.2: Histogram of Grid Population

Notes:

This figure shows the distribution of cell-level population. As with DMSP nighttime lights, there is a large frequency of unpopulated cells.

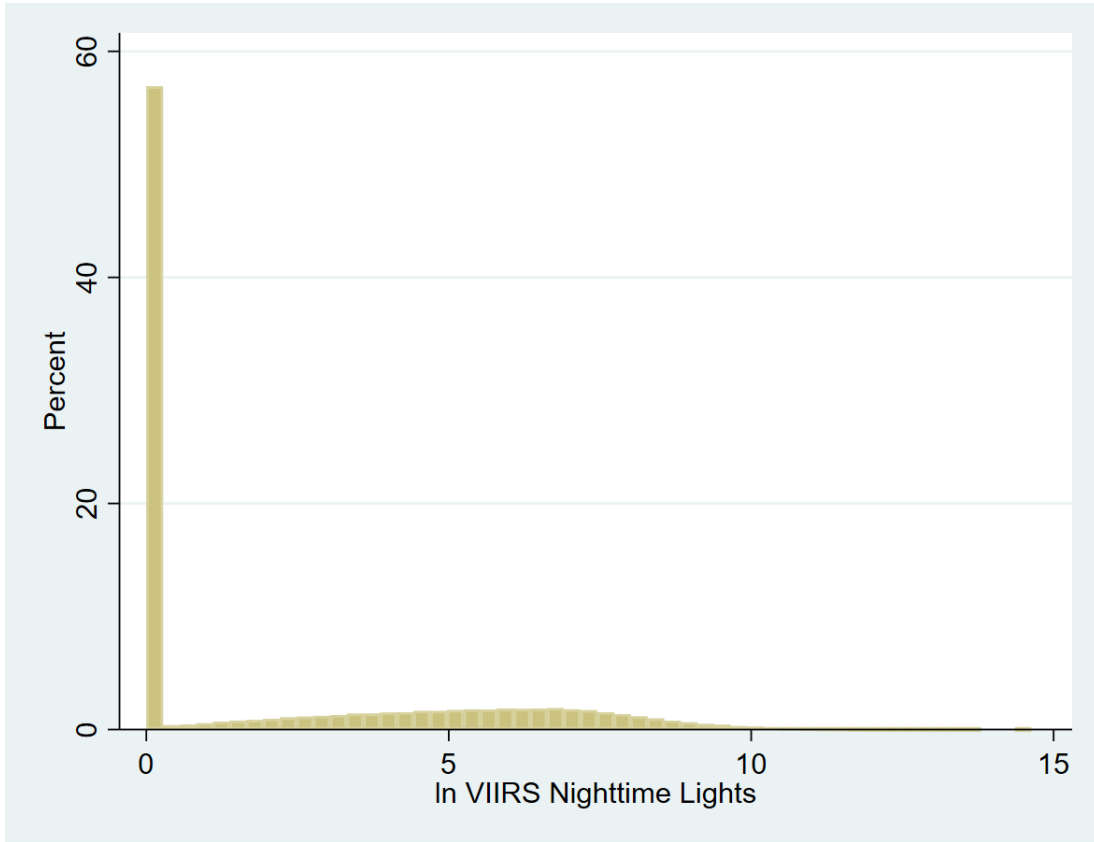


Figure C.3: Histogram of Grid VIIRS Nighttime Lights

Notes:

This figure shows the distribution of VIIRS nighttime lights. Unlike DMSP nighttime lights and population, the majority of observations are not bottom-coded/zeros.

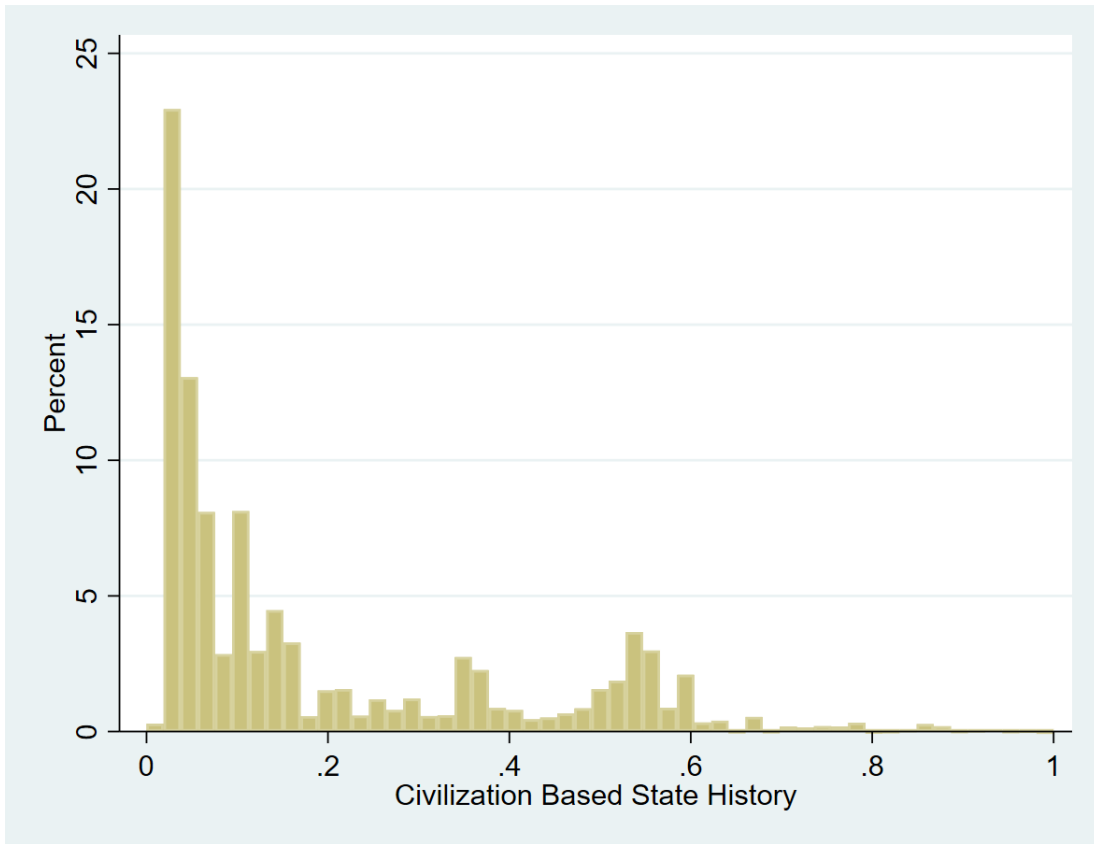


Figure C.4: Histogram of Grid State History

Notes:

This figure shows the distribution of cell-level state history. Unlike DMSP nighttime lights and population, the majority of observations are not bottom-coded/zeros.

Table C.1: Sample Truncation by Outcome Decile

Decile Omitted from Sample:	Bottom (All “0” cells) (1)	Top (2)	Top & Bottom (3)
Panel A. ln DMSP Nighttime Lights, 2010			
State History	2.3476*** (0.5462)	2.6803*** (0.6317)	1.2736*** (0.3958)
Observations	97,181	217,966	72,963
R Sqr.	0.3681	0.4779	0.2232
Panel B. ln Population, 2010			
State History	4.4606*** (0.4699)	3.7657*** (0.4239)	3.9422*** (0.4339)
Observations	217,300	217,965	193,081
R Sqr.	0.7410	0.7585	0.7063
Panel C. ln VIIRS, 2014			
State History	2.5828*** (0.7599)	2.6516*** (0.5719)	1.4101*** (0.5078)
Observations	104,937	217,965	80,718
R Sqr.	0.3254	0.4784	0.2454
Controls (all panels):			
Country FE	Y	Y	Y
All geoclimatic controls	Y	Y	Y

Summary & Notes: This table omits observations by decile of the outcome. Column (1) omits the bottom decile, which accounts for all “0” observations for both nighttime lights and population measures; column (2) omits the top decile, which is a proxy for urban areas; and column (3) omits bottom and top deciles. Statistical significance is determined by country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *.

D Additional Sample Adjustments

Table D.1: Effect by Large (area) Countries

Country:	Russia (1)	Canada (2)	USA (3)	China (4)	All other (5)
Panel A. ln DMSP Nighttime Lights, 2010					
State History	0.7040*** (0.2716)	24.1765*** (2.2864)	20.9960*** (2.2034)	1.4513*** (0.3707)	4.6311*** (0.7250)
Observations	46,179	25,997	18,423	15,337	136,248
R Sqr.	0.5351	0.5789	0.6422	0.6201	0.5871
Panel B. ln Population, 2010					
State History	3.0000*** (0.2030)	26.0365*** (2.5289)	18.3790*** (2.3118)	2.2223*** (0.4017)	4.4464*** (0.5951)
Observations	46,179	25,997	18,423	15,337	136,248
R Sqr.	0.7665	0.6335	0.7386	0.6419	0.7276
Panel C. ln VIIRS Nighttime Lights, 2014					
State History	0.9180*** (0.2698)	21.3653*** (2.2457)	11.8641*** (2.2315)	3.2963*** (0.3634)	4.6540*** (0.7925)
Observations	46,179	25,997	18,423	15,337	136,248
R Sqr.	0.4940	0.4648	0.6406	0.5431	0.5371
Controls:					
Country FE					Y
All geoclimatic controls	Y	Y	Y	Y	Y
Standard Errors:					
Country cluster					Y
HSSW cluster	Y	Y	Y	Y	

Table D.2: Effect by Continent
State History Standardized (mean=0, s.d.=1) Within Continent

Sample:	Old World			New World				
	Africa (1)	Asia (2)	Europe (3)	Old World (4)	South Am. (5)	North Am. (6)	Oceania (7)	New World (8)
Panel A. In Nighttime Lights, 2010								
State History	0.3369*** (0.0480)	0.9699*** (0.2156)	0.1449 (0.1040)	0.7208*** (0.1577)	0.9784*** (0.0837)	0.2573 (0.1770)	0.1477*** (0.0229)	0.5611** (0.2769)
Observations	41,346	51,258	60,050	152,654	25,369	50,941	13,220	89,530
R Sqr.	0.4054	0.5354	0.7136	0.6242	0.4120	0.6658	0.3227	0.5549
Panel B. In Population, 2010								
State History	0.5258*** (0.0874)	0.8531*** (0.1033)	0.3890*** (0.0310)	0.9478*** (0.0762)	0.8148*** (0.0826)	0.1713 (0.1692)	0.2316*** (0.0330)	0.4925* (0.2584)
Observations	41,346	51,258	60,050	152,654	25,369	50,941	13,220	89,530
R Sqr.	0.6849	0.6542	0.8342	0.7693	0.4868	0.7923	0.6698	0.7666
Panel C. In VIIRS, 2014								
State History	0.3544*** (0.0846)	1.0456*** (0.1878)	0.1946* (0.0987)	0.7953*** (0.1657)	0.9931*** (0.0821)	0.1804 (0.1299)	0.1512*** (0.0194)	0.5134** (0.2533)
Observations	41,346	51,258	60,050	152,654	25,369	50,941	13,220	89,530
R Sqr.	0.3943	0.4931	0.6573	0.5816	0.4061	0.6487	0.2269	0.5362
Controls:								
Country FE + All geoclimatic controls	Y	Y	Y	Y	Y	Y	Y	Y

Summary & Notes: Statistical significance is determined by country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *.

E Baseline Results with All Coefficients

Table E.1: Table 1, Panel A with Coefficients of Control Variables

	Dependent Variable: ln Nighttime Lights, 2010					
	(1)	(2)	(3)	(4)	(5)	(6)
State History	5.7222*** (0.9966)	8.0480*** (0.8897)	8.0173*** (0.8573)	3.4938*** (1.0009)	7.9298*** (0.9548)	3.5624*** (0.8074)
Ruggedness (000s of index)			-0.0690*** (0.0207)	-0.0173*** (0.0050)	-0.0659*** (0.0215)	-0.0201*** (0.0051)
index of the stability of malaria transmission, Kiszewski et al. (2004)			-0.0226 (0.0229)	-0.0509*** (0.0123)	-0.0222 (0.0231)	-0.0440*** (0.0105)
tropical moist forest				-0.2037 (0.3282)		-0.2392 (0.2829)
tropical dry forest				0.2838 (0.2425)		0.2253 (0.2140)
temperate broadleaf				1.3612*** (0.2998)		1.1914*** (0.3652)
temperate conifer				0.2986 (0.2799)		0.0725 (0.3067)
boreal forest				-1.0263 (0.6220)		-1.2595* (0.6696)
tropical grassland				-0.0262 (0.2102)		-0.0388 (0.1952)
temperate grassland				0.8631*** (0.2728)		0.8855*** (0.2781)
montane grassland				0.9553*** (0.2714)		0.6692** (0.3015)
tundra				-0.9123* (0.5403)		-1.4013** (0.6802)
Mediterranean forest				1.3232*** (0.4520)		1.0489** (0.4544)
mangroves				-0.1058 (0.3783)		-0.4901 (0.3238)
Average monthly temperature (1960-1990 mon avg)				0.0995*** (0.0295)		0.0972*** (0.0280)
Monthly total precipitation, mm/month (1960-1990 mon avg)				-0.0097*** (0.0021)		-0.0105*** (0.0019)
Length of growing period, days				0.0085*** (0.0016)		0.0082*** (0.0016)
Land suitability as prob that cell is cultivated, Ramankutty (2002)				2.1624*** (0.3471)		2.1704*** (0.3222)
abs(latitude)				0.0289 (0.0300)		0.0283 (0.0290)
Elevation, km above sea level				-0.0992 (0.1709)		0.0737 (0.1983)
Binary variable for if the coast passes inside the grid square					-0.0889 (0.1683)	0.1857 (0.1248)
Distance to the nearest coast, 000s km					-0.4093 (0.6428)	-0.7094** (0.2852)
1(within 25km of natural harbor)					1.8356*** (0.2833)	1.2462*** (0.1780)
1(within 25km of navigable river)					0.8461*** (0.1824)	0.6406*** (0.1442)
1(within 25km of big lake (area _i 5000 sq km))					0.9003*** (0.2975)	0.5475** (0.2206)
Observations	242,184	242,184	242,184	242,184	242,184	242,184
R Sqr.	0.1408	0.3965	0.4068	0.5747	0.4196	0.5857

Summary & Notes: Statistical significance is determined by country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *.

Table E.2: Table 1, Panel B with Coefficients of Control Variables

	Dependent Variable: ln Population, 2010					
	(1)	(2)	(3)	(4)	(5)	(6)
State History	8.4013*** (1.8863)	9.4568*** (1.4901)	9.4860*** (1.4450)	4.2796*** (0.5559)	9.4541*** (1.5534)	4.4186*** (0.4679)
Ruggedness (000s of index)			-0.0180 (0.0276)	0.0195 (0.0143)	-0.0200 (0.0288)	0.0146 (0.0145)
index of the stability of malaria transmission, Kiszewski et al. (2004)			0.0468 (0.0287)	-0.0064 (0.0207)	0.0440 (0.0296)	-0.0034 (0.0188)
tropical moist forest				0.2425 (0.3781)		0.1702 (0.3407)
tropical dry forest				0.5735** (0.2498)		0.5072** (0.2289)
temperate broadleaf				1.9912*** (0.3277)		1.7896*** (0.4042)
temperate conifer				1.0641*** (0.2375)		0.8876** (0.3421)
boreal forest				-0.3693 (0.9537)		-0.6934 (1.0663)
tropical grassland				0.7482** (0.3068)		0.7288** (0.3029)
temperate grassland				1.3925*** (0.2153)		1.3914*** (0.2259)
montane grassland				1.0861*** (0.2866)		0.9880*** (0.2972)
tundra				-0.4480 (0.7471)		-0.9766 (0.9660)
Mediterranean forest				1.8362*** (0.3190)		1.6433*** (0.3786)
mangroves				-0.4485 (0.3413)		-0.3183 (0.3493)
Average monthly temperature (1960-1990 mon avg)				0.1260*** (0.0397)		0.1239*** (0.0377)
Monthly total precipitation, mm/month (1960-1990 mon avg)				-0.0090*** (0.0026)		-0.0091*** (0.0024)
Length of growing period, days				0.0103*** (0.0019)		0.0099*** (0.0017)
Land suitability as prob that cell is cultivated, Ramankutty (2002)				2.0652*** (0.2786)		2.0251*** (0.2347)
abs(latitude)				0.0128 (0.0316)		0.0132 (0.0301)
Elevation, km above sea level				0.2163 (0.2558)		0.3109 (0.2579)
Binary variable for if the coast passes inside the grid square					-0.9217*** (0.2303)	-0.5396** (0.2403)
Distance to the nearest coast, 000s km					-0.4207 (0.8834)	-0.8582** (0.3898)
1(within 25km of natural harbor)					1.5326*** (0.3473)	0.7074*** (0.1479)
1(within 25km of navigable river)					0.6597*** (0.2216)	0.4325*** (0.1458)
1(within 25km of big lake (area _j 5000 sq km))					0.4746* (0.2819)	0.0554 (0.1333)
Observations	242,184	242,184	242,184	242,184	242,184	242,184
R Sqr.	0.2100	0.6147	0.6168	0.7749	0.6238	0.7809

Summary & Notes: Statistical significance is determined by country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *.

Table E.3: Table 1, Panel C with Coefficients of Control Variables

	Dependent Variable: ln VIIRS, 2014					
	(1)	(2)	(3)	(4)	(5)	(6)
State History	5.7467*** (1.0050)	7.9630*** (0.9092)	7.9365*** (0.8898)	3.7159*** (1.0285)	7.9128*** (0.9718)	3.8184*** (0.8424)
Ruggedness (000s of index)			-0.0467** (0.0224)	-0.0043 (0.0062)	-0.0473** (0.0232)	-0.0083 (0.0065)
index of the stability of malaria transmission, Kiszewski et al. (2004)			-0.0235 (0.0219)	-0.0498*** (0.0113)	-0.0253 (0.0229)	-0.0465*** (0.0106)
tropical moist forest				-0.2333 (0.3503)		-0.2914 (0.3161)
tropical dry forest				0.2324 (0.2153)		0.1754 (0.1935)
temperate broadleaf				1.2178*** (0.3404)		1.0436** (0.4032)
temperate conifer				0.2069 (0.2362)		0.0381 (0.2671)
boreal forest				-0.8737* (0.4950)		-1.1483** (0.5778)
tropical grassland				0.0418 (0.1786)		0.0248 (0.1701)
temperate grassland				0.7910*** (0.2346)		0.7934*** (0.2690)
montane grassland				0.7705** (0.2987)		0.6402** (0.3208)
tundra				-0.6602 (0.4175)		-1.1218* (0.5815)
Mediterranean forest				1.1787** (0.4836)		0.9895** (0.4818)
mangroves				-1.0503*** (0.3076)		-1.0317*** (0.2844)
Average monthly temperature (1960-1990 mon avg)				0.1071*** (0.0221)		0.1052*** (0.0217)
Monthly total precipitation, mm/month (1960-1990 mon avg)				-0.0097*** (0.0022)		-0.0099*** (0.0019)
Length of growing period, days				0.0079*** (0.0017)		0.0076*** (0.0016)
Land suitability as prob that cell is cultivated, Ramankutty (2002)				2.1571*** (0.3155)		2.1307*** (0.2938)
abs(latitude)				0.0300 (0.0303)		0.0301 (0.0298)
Elevation, km above sea level				0.0839 (0.1618)		0.1866 (0.1919)
Binary variable for if the coast passes inside the grid square					-0.6400*** (0.1895)	-0.3541** (0.1602)
Distance to the nearest coast, 000s km					-0.4371 (0.5685)	-0.7425** (0.2953)
1(within 25km of natural harbor)					1.3334*** (0.2225)	0.7915*** (0.1404)
1(within 25km of navigable river)					0.7090*** (0.1707)	0.5463*** (0.1169)
1(within 25km of big lake (area _i 5000 sq km))					0.4980*** (0.1848)	0.1893 (0.1326)
Observations	242,184	242,184	242,184	242,184	242,184	242,184
R Sqr.	0.1471	0.3843	0.3895	0.5426	0.3977	0.5499

Summary & Notes: Statistical significance is determined by country-clustered standard errors and denoted at the 1, 5, and 10% levels respectively by ***, **, and *.

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